

IMINCE: an unrestricted factor-analysis-based program for assessing measurement invariance

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Abstract

This paper describes a Windows program for analyzing measurement invariance in two different populations. Factor Analysis is a common way of assessing measurement invariance, and restricted factor analysis is nowadays the most popular method. However, applied researchers have usually found that the theoretical advantages of restricted factor analysis do not always apply in practical situations. For example, when the size of the participants sample is large, as happens in Internet-based questionnaires, the available software for restricted factor analysis might fail to converge to a solution. Our program is based on unrestricted factor analysis and considers the three parameters that define factor invariance: difficulties, discriminations and residual variances. The statistical significance of the tests to evaluate invariance is obtained using Bootstrap resampling procedures. A real example demonstrates the usefulness of the program.

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When we compare members of identifiable groups of individuals for their trait levels, we must assume that the item and test scores that measure the traits have the same meaning in each group. Put more formally, this assumption means that the scores earned by members of different groups are assumed to be on the same measurement scale (Drasgow, 1984). If this assumption is met, the item and test scores are comparable, and the test has ‘measurement invariance’ across the groups. According to the *Standards for Educational and Psychological Testing* (1999, part II), the assessment of measurement invariance is critical to sound testing practice, and so, much discussion and research has been devoted to this topic (see e.g. Reise, Widaman and Pugh, 1993).

Factor Analysis (FA) is one of the most common ways of assessing measurement invariance. The conventional FA approach for checking this issue involves comparing the matrices of item-factor (or test-factor) regression weights of the different groups (see e.g. Jöreskog, 1971). However, this procedure only addresses one aspect of invariance. The general FA model assumes that the regression of an item (or test) score on the factor depends on three parameters: (1) the intercept (i.e. difficulty); (2) the regression weight or factor loading (discrimination); and (3) the residual variance. Strictly speaking, therefore, for two items (or test) scores from two different groups to be comparable, the intercepts, the factor loadings and the residual variances of this item (or test) must be invariant in both groups. Meredith (1993) calls this condition ‘strict factorial invariance’. Following Meredith’s terminology, invariance of the factor loadings would be ‘partial factorial invariance’, whereas invariance in both the intercepts and factor loadings would be ‘strong factorial invariance’.

Historically, the FA assessment of measurement invariance has been addressed from the unrestricted (exploratory) FA model. However, according to Reise *et al.* (1993), the restricted (confirmatory) FA model is more often used nowadays. Restricted FA has important theoretical advantages over unrestricted FA, mainly because: (1) it specifies a structural model which can be rigorously tested, and (2) by choosing a suitable baseline model, we can assess different forms of measurement invariance (partial, strong and strict) by means of hierarchical tests.

Applied researchers, however, have found that the theoretical advantages of restricted FA do not always apply in practical cases. For example, the formal tests of fit used in this model rely on assumptions that are difficult or impossible to fulfil when the variables to be analyzed are item scores (e.g. the assumption that the variables are continuous-unbounded). Furthermore, the standard restricted model assumes that most of the variables are factorially pure (i.e. they only load on one factor and have zero loading values in the remaining factors). In real applications, however, the items tend to have nontrivial secondary loadings on other factors. As some authors noted (Church and Burke, 1994; McCrae, Zonderman, Costa, Bond and Pauonen, 1996), unrestricted FA-based procedures might be more appropriate than the restricted FA approach in most real applications, especially in large multidimensional solutions that do not approach very simple structures. In addition, when the size of the studied sample is large, the available software for restricted FA might fail to converge to a solution. Large participant samples are usually obtained, for example, when it is obtained in Internet- based questionnaires (see, for example, Pasveer and Ellrad, 1998; Buchanan and Smith, 1999; Joinson, 1999).

Because the conventional unrestricted FA approach is mainly descriptive, an important drawback of this model is that decisions are based on arbitrary rules of thumb. To overcome this, several more rigorous procedures have been proposed for assessing item (or test) invariance when an unrestricted FA approach is used. Some of these procedures are inferential and provide standard errors and test statistics, which gives more information and eliminates arbitrariness. However, the relevant procedures are scattered among several journals and, in general, there is no commercial software that implements such procedures (the authors of these procedures usually used ad-hoc routines). Furthermore, all the procedures we revised were only concerned with partial invariance. For these reasons we thought that applied researchers might find useful an unrestricted FA-based general program that allows them to assess the different forms of invariance (partial, strong and strict), and that incorporates a variety of inferential procedures which are not available in commercial programs.

Procedures implemented in IMINCE

IMINCE (an acronym of *Item Measurement INvarianceCE*) is a program written in Visual C 6.0, and is designed to analyze measurement invariance in two populations. Although the

program is particularly suitable for analyses of item scores (either binary or Likert), it can also analyze sums of item scores (parcels) and sets of test scores. In addition, IMINCE is a general purpose program that can be used with any two-group comparison using a Cattell/Cliff-type Procrustes rotation to analyze whole scales. Specifically, the following forms of invariance can be assessed by IMINCE:

a) Invariance of difficulties.

The program tests the general hypothesis that the vector of variable means is the same in the two populations to be compared. This is done using Hotelling's T-square and the corresponding F-ratio. IMINCE also tests the mean differences variable by variable using the univariate t-test. Because the comparisons usually involve large samples, the sizes of the univariate effect (Cohens d') are also reported.

b) Invariance of discriminations (partial invariance).

The discrimination indexes (factor loadings) are computed from the covariance (or the correlation) matrix using three optional methods: Principal Component Analysis, Unweighted Least Squares factor analysis and Unrestricted Maximum Likelihood factor analysis. When the model considers more than one factor (or component), the solution is rotated to show simple structure using Normalized Varimax (Kaiser, 1958) to help the substantive interpretation of the factor solution, and Procrustes (Cliff, 1966) to allow congruence among samples. To test invariance of discrimination indexes, three kinds of tests are implemented in IMINCE: factor congruence, factor discrepancy and approximate confidence intervals for factor loadings. To estimate the discrimination indexes in categorical data, the program allows the so-called 'heuristic approach'. This approach consists of (1) computing the matrix of polychoric correlations between categorical items (tetrachoric correlations in the binary case), and (2) analyzing this matrix by Unweighted Least Squares factor analysis. This approach is simple, deals with large numbers of items and gives similar results to the more theoretically correct approach.

Chan, Ho, Leung, Chan & Yung (1999) proposed a Bootstrap method to evaluate factor invariance in terms of congruence of variables, factors and the overall loading matrix. The method consists of five steps: (1) one sample is taken as the target and one as the replication; (2) the factor solution from the replication sample is rotated against the target using orthogonal Procrustes rotation (Cliff, 1966); (3) empirical congruence indexes between samples are calculated; (4) critical values at α are obtained by Bootstrap resampling; and (5), the observed congruence indices are compared to the critical values at α , and considered as non-statistically significant if they are smaller than the critical value.

Discrepancy of variables, factors and overall loading matrices are evaluated using a similar method. However, the index is based on least-squares measures of fit. In our program, we generalized the overall index proposed by Raykov & Little (1999), so that it is also used for the variables and the factors (as Chang *et al.*, 1999, did for the congruence index). The discrepancy indexes are compared to the critical values at α , and considered as non-statistically significant if they are larger than the critical value.

At the variable level IMINCE also computes approximate confidence intervals for factor loadings. These are bias-corrected percentile intervals obtained from a Bootstrap resampling process (for details see Lambert, Wildt and Durand, 1991). Non-overlapping confidence intervals suggest that a particular variable as a measure of a given factor is not invariant over the two populations of interest.

To compute all the indexes, the user must determine the number of Bootstrap replications from the [500; 5,000] range, and can decide between a 90% or a 95% critical value. It must be noted that usually 1,000 samples are usually recommended in Bootstrap methods (e.g., Efron & Tibshiriani, 1993).

c) Invariance of residual variances.

This form of invariance is assessed variable by variable using bias-corrected percentile intervals obtained from a Bootstrap resampling process. Bootstrap resamples are also drawn from

the [500; 5,000] range, and either 90% or 95% approximate confidence intervals are computed. Nonoverlapping intervals suggest that the residual variances of a particular variable are not invariant over the populations that are compared.

Input and Output

The input and output of IMINCE is illustrated using an empirical example. This is a 10-item Spanish anxiety questionnaire developed by us that uses a 5-point Likert format. The questionnaire was administered to a sample of 707 women and a second sample of 335 men. We aimed to assess the item measurement invariance in the corresponding populations. A model of two factors was expected, and the largest sample was taken as the target sample.

The input consists of two ASCII format files containing participants' scores, the number of participants in each sample, and the number of factors expected in the population. IMINCE default configuration consists of Unweighted Least Squares factor analysis of the covariance matrices, 1,000 Bootstrap samples, and 95% approximate confidence intervals. We used Principal Component Analysis of the covariance matrices and 5,000 Bootstrap samples.

The Output consists of (a) item difficulties, item discriminations and item residual variances for each sample, and (b) the overall, factor and item fit indices described above. Even if the default configuration defines a detailed output, the user can configure the statistics and indices to be reported, that is, stored in the ASCII format file "OUPUT.TXT". The main results are shown in Tables 1, 2 and 3:

- a) Invariance of item difficulties: Hotelling's T-square and univariate t-tests suggest significant differences (see Table 1). However, Cohens d' statistic, which is perhaps more appropriate because the comparisons involved large samples, suggests that there are no substantial differences between populations.

[PLEASE INCLUDE TABLE 1 AROUND HERE]

- b) Invariance of item discriminations: the approximate confidence intervals for factor loadings show overlapping for all the loadings between the populations. However, at the item level there are significant differences in the congruence coefficient of item 1 and in the discrepancy coefficient of item 8 (see Table 2). Because of these significant differences at the item level,

there are also significant differences in the overall congruence and discrepancy indices.

[PLEASE INCLUDE TABLE 2 AROUND HERE]

- c) Invariance of residual variances: the overlapping intervals of all items suggest that the residual variances of items are invariant over the populations compared (see Table 3).

[PLEASE INCLUDE TABLE 3 AROUND HERE]

In a second analysis, we omitted items 1 and 8. Without these two items, IMINCE reported perfect invariance of item difficulties, discriminations and variances. The conclusion of our study was strict factor invariance for items 2, 3, 4, 5, 6, 7, 9 and 10; and no factor invariance for items 1 and 8.

Limitations

We implemented IMINCE for a PC computer using the WINDOWS 95/98/NT operative system. The program uses all the extended RAM memory available in the computer, and the matrices are defined during the execution of the program. This means that there is no clear limit to the maximum number of items that can be analyzed: this depends on the characteristics of the computer that carries out the analyses. The main limitation of IMINCE is the time needed for computing, especially when a larger number of Bootstrap samples is defined. The example in this paper, which in fact involved large samples, was performed on a Pentium III at 866Mhz and 64MB RAM computer. For 5,000 Bootstrap samples, IMINCE needed six minutes and fifteen seconds. However, this time was thirty-nine seconds for 500 Bootstrap samples. When computing polychoric correlation matrices with the standard computers available, the analysis can take a really long time. For 5,000 Bootstrap samples and polychoric correlations, IMINCE needed two hours and twenty-three minutes. In the not-too-distant future, most computers will be able to deal easily with this analysis.

Program availability

A copy of the software, a demo, and a short manual can be obtained at no charge by e-mail (uls@fcep.urv.es).

References

- American Educational Research Association, American Psychological Association , and National Council on Measurement in Education. (1999). Standards for educational and psychological testing. Washington: American Educational Research Association.
- Buchanan, T. & Smith, J.L. (1999). Using Internet for psychological research: Personality testing on the World-Wide Web. British Journal of Psychology, 90, 125-144.
- Chan, W.; Ho, R. M.; Leung, K.; Chan, D. K-S. and Yung, Y-F. (1999). An alternative method for evaluating congruence coefficients with Procrustes rotation: a Bootstrap procedure. Psychological methods, 4, 378-402.
- Church, A.T. and Burke, P.J. (1994). Exploratory and confirmatory tests of the Big Five and Tellegen's three-and-four-dimensional models . Journal of Personality and Social Psychology, 66, 93-114.
- Cliff, N. (1966). Orthogonal rotation to congruence. Psychometrika, 31, 33-42.
- Drasgow, F. (1984). Scrutinizing psychological tests: measurement equivalence and equivalent relations with external variables are central issues. Psychological Bulletin, 95, 134-135.
- Efron, B. and Tibshiriani, R. J. (1993). An introduction to the bootstrap. New York: Chapman & Hall.
- Joinson, A. (1999). Social desirability, anonymity, and Internet-based questionnaires. Behavior Research Methods, Instruments, & Computers, 31, 433-438.
- Jöreskog, K.G. (1971). Simultaneous factor analysis in several populations. Psychometrika, 36, 409-426.
- Kaiser, H.F. (1958). The varimax criterion for analytic rotation in factor analysis. Psychometrika, 23, 187-200.
- Lambert, Z.V., Wildt, A.R. and Durand, R.M. (1991). Approximating confidence intervals for factor loadings. Multivariate Behavioral Research, 26, 421-434.
- McCrae, R.R., Zonderman, A.B., Costa, P.T., Bond, M.H. and Pauonen, S.V. (1996). Evaluating

replicability of factors in the revised NEO personality inventory: confirmatory factor analysis versus Procrustes rotation. Journal of Personality and Social Psychology, 70, 552-566.

Meredith, W. (1993). Measurement invariance, factor analysis and factorial invariance. Psychometrika, 58, 525-543.

Pasveer, K.A. & Ellard, J.H. (1998). The making of a personality inventory: help from the WWW. Behavior Research Methods, Instruments, & Computers, 30, 309-313.

Raykov, T. and Little, T.D. (1999). A Note on Procrustes rotation in exploratory factor analysis: a computer intensive approach to goodness-of-fit evaluation. Educational and Psychological Measurement, 59, 47-57.

Reise, S.P. Widaman, K.F. and Pugh, R.H. (1993). Confirmatory factor analysis and item response theory: two approaches for exploring measurement invariance. Psychological Bulletin, 114, 552-566.

Table 1
Item difficulties, univariate t-tests and Cohen's d statistic

Item number	Target sample	Replication sample	Student's t	Effect Size (Cohens d's)
1	3.42	3.13	4.51**	0.30
2	3.42	3.04	5.25**	0.35
3	3.84	3.82	0.28	0.02
4	2.45	2.42	0.40	0.03
5	2.06	2.12	-0.85	-0.06
6	2.55	2.68	-1.87	-0.12
7	3.07	2.78	3.09**	0.20
8	2.71	2.86	-1.94	-0.13
9	2.88	2.94	-0.72	-0.05
10	2.78	2.73	0.66	0.04

** Significant differences

Table 2
Overall fit congruence and discrepancy indices per item

Item number	Congruence values		Discrepancy values	
	Observed	Critical value at alpha = 0.05	Observed	Critical value at alpha = 0.05
1	0.840**	0.872	0.049	0.055
2	0.607	0.582	0.070	0.092
3	0.993	0.979	0.005	0.026
4	0.991	0.987	0.012	0.025
5	0.974	0.959	0.036	0.058
6	0.992	0.989	0.014	0.020
7	0.998	0.995	0.017	0.035
8	0.992	0.986	0.035**	0.027
9	0.994	0.987	0.009	0.031
10	0.999	0.957	0.018	0.060

** Significant differences

Table 3
Bias-corrected percentile intervals of residual variances per item

Item number	Target sample	Replication sample
1	(0.621; 0.789)	(0.764; 1.085)
2	(0.943; 1.179)	(0.948; 1.246)
3	(0.478; 0.624)	(0.476; 0.679)
4	(0.395; 0.518)	(0.462; 0.687)
5	(0.768; 0.994)	(0.693; 1.113)
6	(0.379; 0.496)	(0.360; 0.554)
7	(0.087; 0.695)	(0.175; 0.680)
8	(0.460; 0.707)	(0.376; 0.680)
9	(0.458; 0.911)	(0.467; 0.976)
10	(0.863; 1.156)	(0.740; 1.070)